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An Examination of the Relationship between Financial Development and Economic Growth in Nigeria: Application of Multivariate Var Framework

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Abstract

This study examined the relationship between financial development and economic growth in Nigeria using annual data for the period 1981-2014. The study employed multivariate VAR framework approach to co-integration to evaluate the long-run relationships between financial development and economic growth. Three financial indicators were used: deposit money bank assets as percentage GDP, ratio of liquid liabilities to GDP and ratio of private sector credit of deposit money banks to GDP. The result found that real gdpc and financial development variables have at least one common stochastic trend driving their relationship. Through VECM granger causality framework, the result found that there is long run unidirectional causality running from economic growth to liquid liability and deposit money bank assets while deposit money bank assets have little significant influence on real gdpc especially at long run. We found feedback effect between private sector credit of deposit money banks and economic growth at the long run. The findings in this study have some important policy implication. To promote the financial system and at the same time promote economic growth, monetary authorities must ensure that banks provide necessary funds to the real sector of the economy.

Key Words: financial development, economic growth, liquid liability, causality, vector error correction

Introduction

The relationship between financial development and economic growth has long been established both at theoretical and empirical levels. However, the emergence of new theories of endogenous growth has indeed renewed interest in the potential role of financial systems in promoting economic growth and development. Greenwood and Jovanovich (1990), Pagano (1993) and King and Levine (1993) had all shown in their studies that financial development does have a positive impact on economic growth through investment, saving, productivity of capital and effective management of information.

There are strands of literature in the study of financial development and economic growth especially in time series studies in the direction of causality. Patrick (1966) explained that finance can lead to economic growth through what he termed the "supply-leading" hypothesis; and equally that economic growth can also stimulate financial development - he called this the "demand following" hypothesis. Ever since the formulation of these hypotheses, empirical conclusions on the direction of causality between financial development and economic growth have remained inconclusive.

In recent years there has been an increase in the application of multivariate VAR model to time-series studies on financial development and economic growth. Prominent among these include Luintel and Khan (1999), Chang and Caudill (2005), Liang and Teng (2006), Ang and Mckibbin (2007), Abu-Badr and Abu-Qarn (2008), Masih et al. (2009), Gries et al. (2009), and Wolde-Rafael (2009). This is because the endogenous growth models have explained that the interaction between financial development and growth often occurs through a number of channels for example through investment, productivity and savings. Therefore, recent empirical works are now exploring some of these channels through the application of multivariate VAR methodology. However, this has still not resolved the issue of causality between financial development and economic growth.

Based on the above, the current study made testable hypothesis that was empirically carried out in the course of this study.

- 1. Ho: $\beta = 0$ There is no long-run significant relationship between economic growth and the financial development indicators (proxied by liquid liability, deposit money bank assets, private sector credit of deposit money banks)
- 2. Ho: $\beta = 0$ There is no causal significant relationship between the economic growth and liquid liability in Nigeria

- 3. Ho: $\beta = 0$ There is no causal significant relationship between deposit money bank assets and economic growth in Nigeria
- 4. Ho: $\beta = 0$ -There is no causal significant relationship between the private sector credit of deposit money banks and economic growth in Nigeria

Theoretical Framework: Financial Development and Economic Growth

There is a growing literature on the impact of financial development on economic growth of the Nigerian economy as a result of the increasing growth in the financial system. According to financial development literature: Patrick (1966), Gurley and Shaw (1955), Goldsmith (1969), Hicks (1969), McKinnon (1973) and Shaw (1973) there is causation that runs from financial development to economic growth implying that financial markets and institutions will increase the financial services. This will definitely lead to high but sustainable real economic growth. Prior to then the relationship between financial development and economic growth has caught the attention of economist such as Schumpeter (1911) who argued that the services provided by financial intermediaries such as mobilizing savings, evaluating projects, managing risk, monitoring managers and facilitating transaction are essential for the technological innovation and economic development of a nation; although the channel and even the direction of causality have remained unresolved in both theory and empirical discuss.

At the forefront of these studies are these researchers: Gurley and Shaw (1955), Goldsmith (1969) and Hicks (1969) who argued that development of a financial system is crucial in stimulating economic growth and under-developed financial systems retard economic growth hence policies aimed at expanding the financial system should be formulated in order to foster growth. McKinnon (1973) and Shaw (1973) have also emphasized on the role of financial intermediaries and financial markets in the growth process. The McKinnon model assumes that investment in a typical developing economy is mostly self-financed hence given its lumpy nature, investment cannot materialize unless sufficient saving is accumulated in the form of bank deposits (McKinnon, 1973). Also, Shaw (1973) further argued that financial intermediaries promote investment and raise output growth through borrowing and lending. This is achieved through mobilization of resources from surplus economic units to deficit economic units. In doing this, they evolved appropriate structures necessary for the intermediation functions which they perform. The result of such financial liberalization, Ang (2007) argued, will lead to increased output growth.

Studies on Financial Development and Economic Growth in Nigeria

Several empirical studies have been carried out to examine the impact of financial development in Nigeria and other developing countries with the ultimate objectives of prescribing measures to enhance financing. Some of the studies will be reviewed,

particularly those with some relevance to Nigeria. Azege (2004) examined the empirical relationship between the level of development by financial intermediaries and growth. The study employed data on aggregate deposit money bank credit over time and gross domestic product to establish that a moderate positive relationship exists between financial development and economic growth. He concluded that the development of financial intermediary institutions in Nigeria is fundamental for overall economic growth.

Nzotta and Okereke (2009) examined financial development and economic development in Nigeria between 1986 and 2007. The study made use of time series data and two stages least squares analytical framework and found that four of the nine variables; lending rates, financial savings ratio, cheques/GDP ratio and the deposit money banks/GDP ratio had a significant relationship with financial development and concluded that the financial system has not sustained an effective financial intermediation, especially credit allocation and a high level of monetization of the economy.

Empirical Evidence from Advanced and Emerging Market

Various academic researchers have examined and documented the link between finance and economic growth in varying dimensions. For example, Wadud (2005) examined the long-run causal relationship between financial development and economic growth for 3 South Asian countries namely India, Pakistan and Bangladesh. The study employed a co-integrated vector autoregressive model to assess the long-run relationship between financial development and economic growth. The results indicate causality between financial development and economic growth but running from financial development to economic growth.

Waqabaca (2004) examined the causal relationship between financial development and growth in Fiji using low frequency data from 1970 to 2000. The study employed unit root test and co integration technique within a vicariate VAR framework. Empirical results suggest a positive relationship between financial development and economic growth for Fiji with causality running from economic growth to financial development. He posited that this outcome is common with countries that have less sophisticated financial systems.

Arestis and Demetriades (1977), using Johansen co-integration on time series analysis for the United states and Germany found insufficient evidence to claim that financial development spurs economic growth. Their analyzed data rather pointed to the direction that real GDP contributes to both banking system and stock market development.

Odiambho (2001) investigated the finance-growth nexus in South Africa using cointegration approach and vector error correction model on monetization ratio; namely the ratio of M2 to GDP and intermediation ratio, the ratio of bank claims on the private sector to GDP against economic growth proxied by real GDP per capital. Their results revealed demand following response between financial development and economic growth and totally discredited the supply leading hypothesis.

Guryay et al (2007), examined the relationship between financial development and economic growth for Northern Cyprus for the period 1986 - 2004 and concluded that there was a negligible positive effect of financial development on economic growth of Northern Cyprus. Rather, the analysis shows that there was evidence of causality from economic growth to the development of financial intermediaries. (King and Levine, 1993 and Levine, Zervos, 1998). Rajan and zingales (1908) using time series analysis (1980-1990) found that financial development has a strong effect on economic growth. Also thngevelu et al, (2004) time series analysis for Australia study represents evidences that financial, markets have causal effect on growth.

Murinde and Eng (1991), Luintel and Khan (1999) argued that a number of endogenous growth models show a two-way relationship financial development and economic growth (Kar and Pentecost 2000). Rousseau and Wachtel (200f) re-examined the core cross country panel result and found that the impact of financial depending on growth is not as strong with more recent data as.it appeared in the original panel studies with data for the period from 1960-19891 and suggested that financial development has a positive effect on growth if not done to excess. Altay and Atgur (2010) advocated bidirectional causality hypothesis. In this study, financial development and economic growth relationship - using VAR Model approach were investigated in Turkey over the period 1970-2006. His empirical findings showed that there was a bidirectional Granger causality relationship between financial development and economic growth in Turkey.

Bailey (2002), studying the relationship between financial sector and economic growth in transition countries was of the view that increased competition in the banking sector (which leads to higher deposit and lower loan rate) has not caused economic growth in Spain province. Stern listed several topics omitted from the survey that are worthy of future research, and financial development was not even mentioned on that list. Future neglect of the role of financial development in economic development is also found in Myers and Seers (1984) book, which is a collection of essays by pioneers of development economics.

Data Description and Methodology

Data Description

This study used annual data covering the period from 1981 to 2014 to investigate the impact of financial development on economic growth.

 Table 1: List of variables and explanations

VARIABLES	YEAR	EXPLANATION & APRIORI EXPECTIONS	SOURCE
Liquid Liability as a percentage of GDP (llgdp)	1981-2014	The ratio of liquid liabilities to GDP measures the size of the financial intermediary system relative to the size of the Nigerian economy and the ability of financial intermediaries to meet unanticipated demand to withdraw deposits by customers (Naceur et al., 2014),	& Levine, 2000.Financial development and structure data set (2014
Real Gross domestic product (RGDP) We expect (+)	1981-2014	Real gross domestic product (GDP) is an inflation-adjusted measure that reflects the value of all goods and it captures the demand and supply of financial activities in the economy.	Development Indicators (World
Deposit Money Bank Asset (dbgdp)	1981-2014	The ratio of deposit money bank assets to GDP captures the overall size of the banking sector relative to the size of the Nigerian economy.	& Levine,
Private credit by deposit money bank as % of gdp (pcrdgdp)	1981-2014	credit to the private sector by deposit money banks (% GDP) which excludes credit issued to the public sector (government, government agencies and public enterprises as well as the credit issued by the monetary authority), The volume of domestic credit to the private sector by deposit money banks relative to the size of the Nigerian economy measures the contribution of financial intermediaries to private sector activities through intermediation.	& Levine, 2000. Financial development and structure data set
Government expenditure % gdp (govgdp)	1981-2014	This variable is included in the study to control for the influence of other components of the Nigerian macro economy.	

Source: Author's Design.

Model specification

Following studies finance-growth literature, it is plausible to form the long-run relationship between *llgdp*, *dbgdp*, *pcrdgdp*, *govxgdp* and *rgdpc*in linear form, with a view of testing the relevant hypothesis as stated in section as follows:

$$rgdpc = f(llgdp, dbgdp, pcrdgdp, govxgdp)$$
 (1)

The above equation can be written in econometric model and in their respective natural log form as thus; The above models can be re-written as econometric model for this study as thus:

$$lnrgdpc_t = \alpha + \beta_1 lnllgdp_t + \beta_2 lndmbgdp_t + \beta_3 lnpcrdgdp_t + \beta_4 rgovxgdp_t + \varepsilon_t$$
 (2)

In the production function $lnrgdpc_t$ is the natural log of income per capita,, $lnllgdp_t$ is the natural log liquidity liability, $lndmbgdp_t$ is the natural log of deposit money bank, $lnpcrdgdp_t$ is natural log private credit to deposit money bank, $lngovexgdp_t$ is the natural log of government expenditure α is the intercept, β_1 to β_4 are the elasticities with respect to change to income per capita. ε_t is the stochastic error term. In line the market fundamentals, we expect all the financial development to have positive signs. However, some studies like Naceur et el (2014) found these variables to have negatively effect in economic growth especially oil exporting countries.

Estimation Procedure

1. Unit root Test

In time series analysis, before running the co integration test the variables must be tested for stationary. For this purpose, we use the conventional ADF tests. Therefore, before applying this test, we determine the order of integration of all variables using unit root tests by testing for null hypothesis $H_0: \beta = 0$ (i.e. β has a unit root), and the alternative hypothesis is $H_1: \beta < 0$. This is to ensure that all the variables are integrated at I(1) to avoid spurious result.

2. Johansen Co integration

This study adopted a dynamic vector autoregressive regression (VAR) which explores co-integration. The essence is to capture the causal dynamics relationship between monetary policy and exchange rate, and at the same time to observe the long run and short dynamics. For instance, given a VAR with possible long run co integration amongst a set of variables.

Therefore, we start with the Johansen co-integration equation which starts with the vector auto regression (VAR) of order p is given by:

$$y_t = \mu + A_1 y_{t-1} + \dots + A_p y_{t-p} + \varepsilon_t$$
 (3)

Where y_t is a $(n \times 1)$ vector of variables under consideration in log form that are integrated at order one- commonly denoted I(1), $n=4A_P$ are the parameters to be estimated, ε_t are the random errors. This (VAR) can be re-written as;

$$\Delta y_t = \mu + \prod y_{t-1} + \sum_{i=1}^{p-1} \Gamma_i \Delta y_{t-i} + \varepsilon_t \tag{4}$$

Where,
$$\Pi = \sum_{i=1}^{p} A_i - 1$$
 and $\Gamma_i = -\sum_{j=i+1}^{p} A_j$ (5)

The above equation is a pure Johansen Co integration test. Gregory and Hansen (1996) noted that the Johansen test is a test for co-integration that allows for more than one co-integration relationship. If the coefficient matrix Π has reduced rankr < n, then there exist $n \times r$ matrices of α and β each with rank r such that

$$\Pi = \alpha \beta' \tag{6}$$

Where r is the number of co-integrating relationship, the element is α is known as the adjustment parameters in the vector error correction model and each column of β is a co integrating vector. It can be shown that, for a given r, the maximum likelihood estimator of β define the combination of y_{t-1} that yield the r largest canonical correlations of Δy with y_{t-1} after correcting for lagged differences and deterministic variables when present. The two different likelihood ratio test of significance of these canonical correlations are the trace test and maximum eigenvalue test, shown in equation 5 and 6 respectively below

$$\lambda_{trace}(r) = -T \sum_{i=r+1}^{n} \ln(1 - \widehat{\lambda_i})$$
 (7)

and

$$\lambda_{max}(r,r+1) = -Tln(1 - \hat{\lambda}_{r+1}) \tag{8}$$

Here, T is the sample size and $\hat{\lambda}_i$ is the i^{th} ordered eigenvalue from the Π matrix in equation 3 or largest canonical correlation. The trace tests the null hypothesis that the number of r co-integrating vector against the alternative hypothesis of n co-integrating vector where nthe number of endogenous variables is. The maximum eigenvalue tests the null hypothesis that there are r co integrating vectors against an alternative of r+1 (see Brooks 2002).

3. Vector Error Correction Model (VECM) and Granger Causality Test

After testing for co-integration among the variables, the long run coefficients of the variables are the estimated. This study uses the Engle and Granger (1987) test augmented by the error correction term for detecting the direction of causality between the variables. The advantage of using vector error correction (VECM) modelling framework in testing for causality is that it allows for the testing of short-run causality through the lagged differenced explanatory variables and for long-run causality through

the lagged ECM term. A statistically significant ECM_{t-1} term represents the long-run causality running from the explanatory variables to the dependent variable. For instance, if two variables are non-stationary, but become stationary after first differencing and are co integrated, the pth-order vector error correction model for the Granger causality test assumes the following equation:

$$\Delta lnX_{t} = \alpha_{10} + \sum_{i=1}^{p_{11}} \theta_{11i} \Delta ln X_{t-1} + \sum_{i=1}^{p_{12}} \partial_{12j} \Delta ln Y_{t-j} + \delta_{13} ECM_{t-1}$$

$$+ u_{1t} \qquad (9)$$

$$\Delta lnY_{t} = \alpha_{20} + \sum_{i=1}^{p_{21}} \theta_{21i} \Delta ln X_{t-1} + \sum_{i=1}^{p_{22}} \partial_{22j} \Delta ln Y_{t-j} + \delta_{23} ECM_{t-1}$$

$$+ u_{2t} \qquad (10)$$

Where θ and $\boldsymbol{\partial}$ are the regression coefficients, u_t is error term and p is lag order of x and yTable 6 indicates that the optimal lag order based on the Akaike information criteria (AIC) is 2. The presence of short-run and long-run causality can be tested. If the estimated coefficients of y in Eq. 2 is statistically significant, then that indicates that the past information of y (e.g economic growth) has a statistically significant power to influence x (Financial development) suggesting that y Granger causes x in the short-run. The long-run causality can be found by testing the significance of the estimated coefficient of ECM_{t-1} (δ_{23}). ECM_{t-1} is the error correction term obtained from the co integration model. The error coefficients δ_{23} indicate the rate at which the co integration model corrects its previous period's disequilibrium or speed of adjustment to restore the long run equilibrium relationship. A negative and significant ecm_{t-1} coefficient implies that any short run movement between the dependant and explanatory variables will converge back to the long run relationship. Indeed, it recovers any long-run information that is partially lost in the system with differenced coefficient. So, that this terms are needed to gain model stability in the long run. Narayan and Smyths (2008).

Data Presentation and Analysis

Our analysis here divided into namely; descriptive statistics and empirical analysis

Empirical Analysis

Stationary test

Table 2: Unit root test

	Augmented Dickey Fuller (ADF)					
Variables	Levels	Levels			Order of Integration	
	t-Stat.	P-value	t-Stat.	P-value	I(1)	
Lngdp	1.117166	0.9968	-3.4857	0.0153	I(1)	
Lndbagdp	-2.48897	0.1274	-3.80848	0.007	I(1)	
Lnllgdp	-2.75774	0.2224	-4.28398	0.01	I(1)	
Lnpcrd	-0.37095	0.5431	-4.49063	0.0013	I(1)	
Lngovexp	-2.04853	0.2658	-7.32732	0	I(1)	

^{**}level of significance at 5% ***level of significant at 1%

Source: Various computation from eview9

The stationary tests were performed first in levels and then in first difference to establish the presence of unit roots and the order of integration in all the variables. The results of the ADF stationary tests for each variable show that the tests fail to reject the presence of unit root for data series in level, indicating that these variables are non-stationary in levels. The first difference results show that these variables are stationary at 1% and 5% significance level (integrated of order one 1(1)). As mentioned in the preceding sections, a linear combination of I (1) series could be I (0) if the series are co integrated. We thus proceed to test for co integration of the time series.

Multivariate Johansen Cointegration-Maximum Likelihood Framework

Table 3: Johansen co-integration result

Hypothesis	Trace Stat	5% critical value	Max.Engen Value	5% critical value		
rgdpc=f(dbagdp,llgdp,pcrd,govexp)						
r=0	86.641**	69.81889	40.05**	33.87687		
r≤1	46.59165	47.85613	28.92354	27.58434		
r≤2	17.6681	29.79707	9.46878	21.13162		
r≤3	8.199323	15.49471	8.129614	14.2646		
r≤4	0.069709	3.841466	0.069709	3.841466		

^{*}level of significance at 10% **level of significance at 5% ***level significance at 1%

Source: various computation from eview9

The result of the co integration test, based on the Johnson co integration approach are presented in table 5. The aim is to establish whether long-run relationship exists among the variables of interest. The first step was to establish the lag using akaike information

criterion. Then, cointegration was tested on the long run relationship between the dependent variable (real gdpc) and independent variables (financial development variables) while controlling for the influence government expenditure. The table indicates that the test failed to accept the null hypothesis of no cointegration at 5% level of significance. Both the trace and Maximum Eigen value suggest that there is a common stochastic trend and as such the number of free random walks has been reduced by one. This implies that an equilibrium relationship exists among the co integrating variables. In addition, no matter the fluctuation in the short run, these variables have the tendency to return to this equilibrium path in the long run.

Normalised cointegartion Equation:

$$\Delta rgdpc = -20.94 - 0.972dmb - 0.5057ll + 2.125pcrd - 1.102govexp + \varepsilon_t$$
(11)
$$[1.9768] \quad [2.0068] \quad [-3.805] \quad [6.3062]$$

Equation 11 represents the normalised co integration equation, while the values in the bracket are the t-statistics. The equation gives the long run impact of financial development on economic growth captured by various variables deposit money bank, and liquidity liability which contributed negatively significant to economic growth in Nigeria while and private credit by deposit money bank has a positive and significant impact on the real gdpc. Indeed, all the variables do not agreement with the apriori expectations, except PCRD and the figures in bracket are the t-statistics.

The study found the coefficient of DMB to be negative and statistically significant at 10% level. With coefficient of 0.9872, 1% decrease in deposit money bank will cause the overall economic growth increase by 0.9872% in the long run. Also, the study found the coefficient of liquidity liability to be negative and statistically significant at 5% level. With coefficient of 0.5057, 1% decrease in liquidity liability will cause the overall economic growth increase by 0.5057% in the long run. Furthermore, study found the coefficient of PCRD to be positive and statistically significant at 5% level. With coefficient of 2.125, 1% decrease in PCRD will cause the overall economic growth increase by 2.125% in the long run. Controlling for the influence of government expenditure, the study found the coefficient of government expenditure to be negative and statistically significant at 10% level. With coefficient of 1.12, 1% decrease in government expenditure will cause the overall economic growth increase by 1.12% in the long run.

Since the presence of co integration among variables means that causality must run from at least one direction, therefore, we apply error correction model, in company of variance decomposition and impulse response for more robust analysis

Causality Test

Table 4: Long run and short run causality estimates.

		VECM Granger Causality					
Type of Causalit	у						
		Short run	Excluded variables				Long run
Dep. Var		D(LGDP)	D(LDBAGDP)	D(LLLGDP)	D(LPCRD)	D(LGOVEXP)	ECT
D(LGDP)	Chi-sq		1.209608	0.290412	2.793513	3.442901	-0.08353
	P- value		(0.5462)	(0.8648)	(0.2474)	(0.1788)	[-2.1708]
D(LDBAGDP)	Chi-sq	0.178756		5.882083	0.694323	0.971385	0.140703
	P- value	(0.9145)		(0.0528)	(0.7067)	(0.6153)	[0.5895]
D(LLLGDP)	Chi-sq	0.324074	1.895006		0.423421	0.516865	0.09818
	P- value	(0.8504)	(0.3877)		(0.8092)	(0.7723)	[0.4170]
D(LPCRD)	Chi-sq	1.138886	3.202484	8.310415		3.878337	-0.17583
	P-						
	value	(0.5658)	(0.2016)	(0.0157)		(0.1438)	[-2.8007]
D(LGOVEXP)	Chi-sq	5.299467	0.10225	2.455385	8.453953		-0.65345
	Chi-sq	(0.0707)	(0.9502)	(0.293)	(0.0146)		[-2.7686]

Note: *, **, and *** indicate significance at 10%, 5% and 1%, respectively t-statistics in [] and P-values in ()

Source: various computations from eview9

This study used the Granger causality test augmented by the error correction term for detecting the direction of causality between the variables. The optimal lag order selected based on the Akaike information Criteria (AIC) is 2. The VECM Granger causality divides causality results into long run as well as the short run. The results regarding the VECM Granger causality test are reported in Table 6. The empirical results suggest that ECT_{t-1} has negative sign and statistical significant in economic growth, PCRD and govexp. This implies that there is bidirectional causality between economic growth, PCRD and govexprespectively in the long run. Bidirectional causality between economic growth, PCRD and govexp indicate that they are complementary.

A number of causal interactions exist in the short run. The results in Table 4 show a unidirectional causality running from government expenditure to economic growth, from economic growth to government expenditure. There are no other causalities reported in table 6 indicating that financial activities in Nigeria is yet to be exploited. In sum, the coefficients of ECM (-1) in table 4 is negative and significant at 5% level.

The coefficients suggest that approximately 8% of the short-run disequilibrium is corrected in the long run within one-year period.

Discussion of Findings

The endogenous growth models suggest that financial development influences economic growth through a number of channels particularly through investment, saving, and productivity. In agreement to this, we examined financial development and economic growth using various channels. Using the individual financial indicators, the results are as follows; The Trace and Max Engen statistics are significant at 5% level suggesting that there is a common stochastic trend and as such the number of free random walks has been reduced by one. Therefore, real gdpc and financial development variables have at least one common stochastic trend driving the relationship between them. The idea here is that the null hypothesis of no long run significant relationship of economic growth and financial development variables is rejected.

In the second hypothesis, VECM granger causality test augmented by the error correction term was used to establish the direction of causality between economic growth and financial development variables. The result revealed that there is long run unidirectional causality running from economic growth to liquid liability at 5% level of significance. This result followed the findings of King and Levine, 1993 and Levine, Zervos, 1998 who posited that there was evidence of causality running from economic growth to the development of financial intermediaries. Whereas, in the short run, causalities do not run from any direction. The intuition here is that real gdpc causes liquid liability in Nigeria. Hence, we reject the null hypothesis of no causal significant relationship between liquid liability and real gdpc in Nigeria.

The third hypothesis centred on causal relationship between the deposit money bank and economic growth in Nigeria using augmented error correction term. Our result revealed that there is long run unidirectional causality running economic growth and deposit money bank at 5% level of significance as expected. We also found that change in deposit money bank has little significant influence on real gdpc, especially at long. Therefore, we reject the null hypothesis of no causal significant relationship between deposit money bank and real gdpc in long run and vice versa for the short run.

Turning to Hypothesis 4, our results suggested that ecm_{t-1} has negative sign and statistical significant in the long run indicating that causality is running from real gdpc to PCRD at 5% level of significant. There is also, a noticeable causality running PCRD to economic growth in the long run. It means that the presence of bi-directional causality hypothesis, indicating the combination of supply-leading and demand following hypotheses. This result is in consistence with the finding from (Greenwood & Jovanovic, 1990; Berthelemy & Vardoulakis, 1996; Blackburn & Hung, 1998; and Harrison) who postulated that financial development gradually induces economic growth and this, in turn, causes feedback and induces further financial development. It

is also consistent with endogenous growth models as explained by Murind and Eng (1991), Luintel and Khan (1999) and (Kar and Pentecost 2000). Therefore, we reject the null hypothesis that there is no causal significant relationship between private sector credit and economic growth in Nigeria

Policy Recommendations

In Nigeria, for the financial system to clearly promote economic growth, monetary authorities must ensure that banks provide necessary funds to the real sector of the economy. At present, there is a weak link between the real sector and the financial system, an indication that the majority of banks' loans are channelled to unproductive sector of the economy. Monetary authorities therefore must pursue appropriate policies that will increase the level of financial intermediation, achieve positive interest rate and increase the level of investment.

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