The impact of smallholder irrigation on household welfare: The case of Tugela Ferry irrigation scheme in KwaZulu-Natal, South Africa

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ABSTRACT

The potential of smallholder irrigated agriculture to enhance food security and alleviate rural poverty has led the South African Government to prioritise and invest significantly in irrigation establishment, rehabilitation and revitalisation. The question addressed in this study pertains to the extent to which smallholder irrigation has been able to reduce poverty in the rural communities to justify this investment. Using a sample of 251 farmers, this study found that factors such as land size, perceived soil fertility, household size, and access to support services were significant predictors of irrigation participation. The results from the treatment effect model indicated that access to irrigation plays a positive role in the welfare of rural households, with irrigators spending about ZAR 2,000 per adult equivalent on consumption more than the non-irrigators. The study, therefore, concluded that government investments in smallholder irrigation for poverty reduction are justified. The other factors that influenced household consumption were off-farm income, land size, livestock size, education level, family size and access to support services and infrastructure. The study recommends that investments in smallholder irrigation continue for poverty reduction, and that priority should also be on finding other feasible rural micro-projects and development initiatives to complement smallholder irrigation and significantly reduce rural poverty.

Keywords: smallholder irrigation, poverty, food security, treatment effect model, Foster Greer Thorbecke (FGT) poverty measures

INTRODUCTION

Poverty reduction and ensuring household food security are important policy goals in developing countries, particularly in sub-Saharan Africa. Several authors agree that reaching the Millennium Development Goal (MDG) of halving poverty and hunger by 2015 in this region requires giving high priority to smallholder agriculture (Smith, 2004; Matshe, 2009; Tshuma, 2012). A general consensus is that smallholder irrigation remains a feasible and key strategy for achieving improved agricultural production, household food security and rural poverty reduction in the developing world (Kumar, 2003; Lipton et al., 2003; Hussain and Hanjra, 2004; Gebregziabher et al., 2009; Bacha et al., 2011). Although irrigation development comes at a cost, and may have negative environmental and health consequences such as increased water logging, salinization and water-borne diseases, it is one of the most important factors in increasing crop productivity and improving overall agricultural performance (Hussain and Wijerathna, 2004).

Access to irrigation increases the area under cultivation and crop intensity, and decreases crop losses (Namara et al., 2010). Moreover, it leads to poverty reduction by expanding opportunities for higher and more stable incomes, and by increasing prospects for multiple cropping and crop diversification (Hussain and Wijerathna, 2004). The potential of irrigated agriculture in enhancing food security and alleviating poverty has led the South African Government to prioritise irrigation development (Denison and Manona, 2007; Van Averbeke et al., 2011). The establishment, rehabilitation and revitalisation of smallholder irrigation schemes were made possible through the investment of large amounts of public resources (Denison and Manona, 2007). Shah et al. (2002) estimated the public investments in smallholder irrigation at ZAR 2 billion (ZAR 40 000/ha). In fact, smallholder irrigation schemes continue to be a major budget item on many developmental and district municipality financial plans (Denison and Manona, 2007).

Irrigation farming is imperative in South Africa as rain-fed crop production is inherently risky due to unreliable rainfall and frequent droughts (Cousins, 2012). South Africa is generally dry, with over 60% of the country receiving less than 500 mm of rain per annum on average, and with only 10% receiving more than 750 mm (World Bank 1994 cited in Cousins, 2012). The importance of irrigation farming in South Africa is underscored by the fact that the irrigated 8% of land under crop production contributes almost 30% of total agricultural production (Backeberg, 2006; NDA, 2007; Hope et al., 2008). Smallholder irrigation accounts for about 0.1 million hectares (about 8%) of total irrigated land in South Africa (Tlou et al., 2006; NDA, 2007; Van Averbeke et al., 2011).

Although smallholder irrigation accounts for a small proportion of irrigated area in South Africa, it is important and has generated national public interest in recent years (Denison and Manona, 2007). The importance arises primarily from its location in the rural areas, where poverty and food insecurity are concentrated (Perret, 2002; Sishuta, 2005; Vink and Van Rooyen, 2009). Poverty alleviation and ensuring household food security in rural areas are major objectives for the establishment of smallholder irrigation in South Africa (Denison and Manona, 2007). Furthermore, smallholder irrigation has the potential to create employment in these underdeveloped rural areas, both directly and indirectly through forward and backward linkages (Van Averbeke et al., 2011).

However, many researchers have reported that, despite its potential, smallholder irrigation has failed to meet the rural development and poverty reduction objectives in South Africa.
According to Bembridge (2000), the performance and welfare impact of smallholder irrigation schemes has been poor, and falls far short of the expectations of many stakeholders. Infrastructure deficiencies and poor institutional support have been cited as the major reasons for the failure of smallholder irrigation in South Africa (Machethe et al., 2004; Tlou et al., 2006; Van Averbeke et al., 2011; Fanadzo, 2012). However, this claim of failure has not been based on in-depth and systematic impact evaluations.

Most of the previous studies on the impact of smallholder irrigation schemes in South Africa (Hope et al., 2008; Fanadzo et al., 2010; Van Averbeke, 2012) have been descriptive in nature and have not included any systematic quantitative evaluations. For example, Van Averbeke (2012) used observations of cropping intensity as a performance and/or impact indicator, while Hope et al. (2008) relied on gross margin analysis. Although these evaluations are relevant and make up an important part of the measurements of poverty impacts of smallholder irrigation, they are not complete, as they do not evaluate direct irrigation impact on household welfare. This study aimed, therefore, to provide empirical evidence and systematic quantitative analysis of the impact of smallholder irrigation on household welfare using the Tugela Ferry irrigation scheme as an empirical example. The study was undertaken as part of a project (K5/2176) initiated, managed and funded by the Water Research Commission (WRC) entitled, ‘Empowerment of women in rural areas through water use security and agricultural skills training for gender equity and poverty reduction in KwaZulu-Natal Province’.

**METHODS**

**Study area description**

The Tugela Ferry irrigation scheme is located in Msinga Local Municipality in the Mzinyathi District in the KwaZulu-Natal Province of South Africa. It is situated 120 km north of Pietermaritzburg, the provincial capital. Figure 1 shows the location of the Tugela Ferry area in the Msinga local municipality. The location of the Msinga Local Municipality in South Africa and in the KwaZulu-Natal Province is also shown on the maps.

Msinga is characterised by high poverty levels as it has few economic resources and little economic activity (Dearlove, 2007). Moreover, the area is characterised by high unemployment rates leading many of its inhabitants to be involved in subsistence and informal activities. Almost half of the households earn less than ZAR800 per month (Dearlove, 2007). Subsistence agriculture plays a significant role in the welfare of the poor in the area, and it is mainly practised in areas adjacent to the Tugela and Mooi rivers.

Rain-fed crop production is challenging in Msinga because the area is both hot and dry. Msinga is situated in a dry to semi-arid zone with mean rainfall of 600–700 mm/a and very high summer temperatures of up to 44°C (Cousins, 2012). The area is characterised by frequent droughts, making irrigation the main mode of household food production (Cousins, 2012). Irrigation farming, therefore, offers many of the rural people an opportunity to increase their incomes and participate in the local economy. The Tugela Ferry and Mooi irrigation schemes play an important role in the local economy of Msinga as sources of food, employment and market for agricultural inputs.
The Tugela Ferry irrigation scheme is located on both banks of the Tugela River. Figure 2 shows a sketch map of the Tugela Ferry irrigation scheme. The scheme was planned and constructed by the Natal Native Trust between 1898 and 1902 to improve household welfare, and has been operational ever since (Cousins, 2012). It consists of 7 blocks of irrigable land covering 837 ha, of which approximately 540 ha is flood-irrigated, and is one of the largest smallholder irrigation schemes in the province (Cousins, 2012; Fanadzo, 2012). Water is drawn from the Tugela River in several ways. A canal of 31 km is used to draw water under gravity for Blocks 1–3 and 5; while a diesel pump is used for Block 4B and electric pumps for Blocks 4A and 7. Initially, all the blocks got water from the main canal, but water shortages have meant that only 4 blocks benefit from the gravity-fed canal while other blocks use pumps. The initial canal infrastructure has deteriorated such that it is difficult to supply the four gravity-fed blocks. Irrigators take turns in diverting irrigation water to their plots at least once per week.

A total of about 1 500 irrigators participate in the irrigation scheme growing various crops such as maize, cabbage, potatoes, tomatoes, onions, beans, beetroot, spinach and butternut. The land in the scheme is owned by 3 tribal authorities (Mthembu: Block 1–5; Mabaso: 7A; and Bomvu: 7B). The ownership of Block 6 is uncertain, and it is currently not operating due to tribal conflicts regarding ownership. Farmers in the irrigation scheme were initially allocated 2 plots of 0.1 ha each. Over time, some farmers have managed to acquire more plots through leasing or borrowing from neighbours. The access to land is through the tribal authorities, who allocate land to households. Selling of land is not permissible and land is returned to the tribal authority for re-allocation if the owner is no longer using it.

Sampling techniques and data collection tools

Data were collected over a period of 3 weeks in November 2012 by trained enumerators who were Zulu speaking, the local language. A stratified random sampling technique was used to select the survey respondents. Households were categorised into 2 strata: irrigation participants and non-participants. A list of the irrigation farmers was obtained from extension officers, and farmers were further stratified according to their blocks on the scheme. From these sub-strata, a simple random sample was drawn to obtain 186 irrigating farmers such that every block was represented by at least 10% in the final sample of irrigators. A sample of 186 irrigating households was considered appropriate as it represents above 10% of the 1 500 irrigating farmers in the Tugela Ferry irrigation scheme. There was no list for non-irrigators, therefore, the non-irrigators that were interviewed were identified during the survey. A sample of 70 non-irrigators was selected and interviewed to serve as a proxy for the control or counterfactual group. Since the population is relatively homogenous, this sample size was considered large enough to provide a reliable counterfactual.

Primary data were collected using structured questionnaires and focus group discussions. Questionnaires were pre-tested and modified accordingly before being administered. Information on basic characteristics of household heads such as sex, age, marital status and education level was collected using the questionnaire. The questionnaire also included measures of household wealth such as household assets, livestock, and type of houses; agricultural production activities; and household expenditure, income amounts and sources. The same questionnaire was used for both irrigators and non-irrigators, but with extra sections to cover specific questions related to the irrigation activities. The approach adopted in the survey is in line with Jalal and Ravallion (2003), who suggested that in project impact evaluations it is important that the same questionnaire be administered to both groups, and that project participants and non-participants are from the same economic environment.

Analytical methods

The study employed both descriptive and econometric techniques. Descriptive analysis was performed using the t-test for continuous variables and χ2 test for categorical variables. The Foster Greer Thorbecke (FGT) poverty indices were also used to give a summary of the incidence, depth and severity of poverty in the study area. Econometric analysis employed the treatment procedure to identify the impact of smallholder irrigation on household welfare from among other possible factors influencing household welfare. Propensity score matching (PSM) was used as a robust check on the results of the treatment effect regression model.

The Foster, Greer and Thorbecke (FGT) poverty indices

The Foster, Greer and Thorbecke (FGT) poverty measures were calculated to examine the incidence, depth and severity of poverty among irrigators and non-irrigators as below:

\[ P_\alpha = \frac{1}{n} \sum_{i=1}^{n} \left( \frac{1}{2} \frac{Y_i - \bar{Y}}{\sigma} \right)^\alpha \]

(1)
The poverty aversion parameter is a non-negative parameter indicating the degree of sensitivity of the poverty measure to inequality among the poor. The incidence of poverty (head-count index), estimated when \( \alpha = 0 \), measures the share of the population below the poverty line. The poverty depth index (poverty gap), estimated when \( \alpha = 1 \), captures information regarding how far off households are from the poverty line. The poverty severity index (poverty gap square), estimated when \( \alpha = 2 \), takes into account not only the distance separating the poor from the poverty line (the poverty gap), but also the inequality among the poor.

A larger \( \alpha \) value gives greater emphasis to the poorest of the poor, indicating greater sensitivity of the poverty measure to inequality among the poor (Foster et al., 1984; Namara et al., 2008). It is also important to highlight that \( z \) - \( Yi \) takes the value of zero when \( Yi > z \) in the above equation. The FGT poverty measures were calculated using Distributive Analysis Stata Package (DASP) version 2.2, an ado Stata file created by Araar and Duclos (2012). Based on the minimum per capita adult-equivalent calorific intake (at 2 261 kcal/day) (Frye, 2005; Stats SA, 2007), a figure of ZAR5 277 per adult equivalent per annum was used as the poverty line. This figure was taken from that reported by Stats SA (2007), and then adjusted using the consumer price index (CPI) so that it reflects the current purchasing power of the Rand (ZAR). The adult equivalents were calculated using the scales recommended by Smith and Subandoro (2007), which take age and gender into consideration to adjust household sizes to reflect their energy requirements.

The treatment effect model

The major econometric problem in evaluating project impacts is selection bias (Maddala, 1983). Smallholder irrigation usually purposively targets the poor, who are presumably more likely to be poor without access to irrigation (Baker, 2000). Therefore, it is expected that irrigation participants would have had far less consumption expenditure in the absence of the irrigation project (Baker, 2000), and sample selection bias arises due to this self-selection. Using OLS to estimate the impact of irrigation participation on household welfare when there is selection bias produces biased and inconsistent estimates (Heckman, 1979). The welfare difference between irrigators and non-irrigators, therefore, cannot be attributed to access to irrigation so long as selection bias exists (Bacha et al., 2011).

Heckman (1979)’s basic model of selectivity has been widely applied in evaluating programme benefits (or treatment effects) (Maddala, 1983; Greene, 2003). The model corrects for the selection bias that arises from unobservable factors by estimating 2 equations: the selection (participation) equation and the response (outcome) equation. The discussion of the treatment effect model below derives from mainly Heckman (1979), Maddala (1983) and Greene (2003). Interested readers may consult these texts for further details.

The irrigation participation equation, estimated using the probit model, assumes that irrigation participation is a linear function of the exogenous covariates \( z_i \) and the residual error \( u_i \). Specifically, the irrigation participation model is specified as follows:

\[
P^* = \gamma z_i + u_i,\]

\[
P_i = 1 \text{ if } P^*_i > 0, 0 = \text{ otherwise} \tag{2}
\]

where:
- \( P^* \) is the latent endogenous variable such that \( P^*_i \) takes a value of 1 when \( P^*_i \) is greater than zero
- \( z_i \) is a vector of household characteristics that influence household’s access to irrigation
- \( \gamma \) are the coefficients to be estimated
- \( u_i \) is the residual term

The participation equation is used to generate a selection variable (inverse Mills ratio) which, when included in the OLS model, makes the estimates unbiased and consistent. The second stage of the treatment effect regression procedure involves adding the inverse Mills ratio to the response equation and estimating the equation using OLS as follows:

\[
Y_i = \beta x_i + \delta R_i + \lambda \gamma z_i + \epsilon_i \tag{3}
\]

where:
- \( Y_i \) is total household consumption expenditure per adult equivalent
- \( x_i \) is a vector of household characteristics
- \( R_i \) indicates whether a household is an irrigation beneficiary or non-beneficiary
- \( \lambda \) is the inverse Mills ratio
- \( \epsilon_i \) is the error term
- \( \beta \) and \( \delta \) are parameters that are to be estimated.

\( Y_i \) was a continuous variable measuring total household food and non-food expenses incurred by households in a year. The food items produced and consumed by the household were converted to their market values using average local prices.

Although the impact coefficient \( \delta \) in Eq. (3) is unbiased due to the inclusion of the selectivity term, it is inefficient as the disturbance term \( \epsilon_i \) is heteroscedastic (Greene, 2003). However, the Stata software package, which was used in this study, automatically corrects for that bias in standard errors. An assumption required to guarantee reliable estimates of the outcome equation is the existence of at least one additional regressor (with a non-zero coefficient) in the selection equation which has no direct effect on the outcome (Blundell and Costa-Dias, 2000; Heckman and Vytlacil, 2005). Including the same number of variables in the selection and outcome equations would lead to the multicollinearity problem in the outcome equation which results in very imprecise estimates (Sartori, 2003). Therefore, the consumption expenditure equation included one variable less than the irrigation participation equation for model identification purposes.

In terms of the estimation procedure, both the 2-step estimator (described above) and maximum likelihood have been used (Maddala, 1983). Compared to the maximum likelihood method, 2-step estimation procedure is less efficient (Puhani, 2000). The full information maximum likelihood estimation procedure, therefore, was reported in this study. The next section presents the empirical results and their discussions.

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RESULTS AND DISCUSSION

The results of data analysis are presented in the following sections. The next section presents descriptive statistics from both t-tests and χ²-square tests. The results from FGT poverty analysis are then presented, followed by the results from the econometric model.

Descriptive statistics

The total sample size was 256, comprising of 70 non-irrigators and 186 irrigators. However, 5 households were discarded in the final sample for analysis due to missing information. The results of descriptive analyses are presented in Tables 1 and 2.

TABLE 1
Comparison of continuous variables between irrigators and non-irrigators

<table>
<thead>
<tr>
<th>Variable definition</th>
<th>Non-irrigators (N=66)</th>
<th>Irrigators (N=185)</th>
<th>T-test</th>
</tr>
</thead>
<tbody>
<tr>
<td>Mean</td>
<td>St. Dev.</td>
<td>Mean</td>
<td>St. Dev.</td>
</tr>
<tr>
<td>Household age (years)</td>
<td>58.55</td>
<td>16.08</td>
<td>56.88</td>
</tr>
<tr>
<td>Household size in numbers</td>
<td>5.98</td>
<td>3.14</td>
<td>6.30</td>
</tr>
<tr>
<td>Household size in adult equivalents</td>
<td>4.34</td>
<td>2.24</td>
<td>4.84</td>
</tr>
<tr>
<td>Education level in years</td>
<td>2.48</td>
<td>3.85</td>
<td>2.32</td>
</tr>
<tr>
<td>Land size (ha)</td>
<td>0.59</td>
<td>0.73</td>
<td>0.24</td>
</tr>
<tr>
<td>Value of assets (Rand)</td>
<td>66 950.38</td>
<td>114 753</td>
<td>135 186.8</td>
</tr>
<tr>
<td>Livestock units in TLU</td>
<td>3.18</td>
<td>4.99</td>
<td>3.22</td>
</tr>
<tr>
<td>Off-farm income</td>
<td>42 332.84</td>
<td>28 465.19</td>
<td>36 333.66</td>
</tr>
<tr>
<td>Farm income</td>
<td>321.46</td>
<td>806.98</td>
<td>15 341.31</td>
</tr>
<tr>
<td>Food expenditure per adult equivalent per annum</td>
<td>3 484.42</td>
<td>1999.79</td>
<td>4 631.23</td>
</tr>
<tr>
<td>Non-food expenditure per adult equivalent per annum</td>
<td>642.11</td>
<td>879.48</td>
<td>1 345.14</td>
</tr>
<tr>
<td>Total expenditure per adult equivalent per annum</td>
<td>4 126.53</td>
<td>2 368.14</td>
<td>5 976.37</td>
</tr>
<tr>
<td>Share of food expenditure (%)</td>
<td>0.87</td>
<td>0.14</td>
<td>0.79</td>
</tr>
</tbody>
</table>

***significant at 1%; **significant at 5%; *significant at 10% significance levels.

Source: Household survey (2012)

TABLE 2
Household characteristics frequencies between irrigators and non-irrigators

<table>
<thead>
<tr>
<th>Variable definition</th>
<th>Categories</th>
<th>Non-irrigators (%) (N=66)</th>
<th>Irrigators (%) (N=185)</th>
<th>χ² test</th>
</tr>
</thead>
<tbody>
<tr>
<td>Household head sex</td>
<td>0=female</td>
<td>65.15</td>
<td>66.84</td>
<td>0.063</td>
</tr>
<tr>
<td>1=male</td>
<td>34.85</td>
<td>33.16</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Household marital status</td>
<td>1=single</td>
<td>13.64</td>
<td>12.30</td>
<td>1.06</td>
</tr>
<tr>
<td>2=married</td>
<td>50.00</td>
<td>57.22</td>
<td></td>
<td></td>
</tr>
<tr>
<td>3=divorced</td>
<td>4.55</td>
<td>3.74</td>
<td></td>
<td></td>
</tr>
<tr>
<td>4=widowed</td>
<td>31.82</td>
<td>26.74</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Household main religion</td>
<td>0=no religion</td>
<td>7.58</td>
<td>5.38</td>
<td>1.49</td>
</tr>
<tr>
<td>1=traditional</td>
<td>36.36</td>
<td>35.48</td>
<td></td>
<td></td>
</tr>
<tr>
<td>2=Christian</td>
<td>56.06</td>
<td>57.53</td>
<td></td>
<td></td>
</tr>
<tr>
<td>3=Muslim</td>
<td>0.00</td>
<td>1.61</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Access to extension services</td>
<td>0=no</td>
<td>64.62</td>
<td>29.19</td>
<td>25.52***</td>
</tr>
<tr>
<td>1=yes</td>
<td>35.38</td>
<td>70.81</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Access to credit</td>
<td>0=no</td>
<td>76.92</td>
<td>34.22</td>
<td>2.7852*</td>
</tr>
<tr>
<td>1=yes</td>
<td>23.08</td>
<td>65.78</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Access to road</td>
<td>0=no</td>
<td>53.73</td>
<td>17.11</td>
<td>22.48***</td>
</tr>
<tr>
<td>1=yes</td>
<td>46.27</td>
<td>82.89</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Training</td>
<td>0=no</td>
<td>90.77</td>
<td>88.17</td>
<td>0.3278</td>
</tr>
<tr>
<td>1=yes</td>
<td>9.23</td>
<td>11.83</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Perception of land quality</td>
<td>0=bad</td>
<td>30.77</td>
<td>12.85</td>
<td>25.53***</td>
</tr>
<tr>
<td>1=good</td>
<td>32.31</td>
<td>44.18</td>
<td></td>
<td></td>
</tr>
<tr>
<td>2=very good</td>
<td>36.92</td>
<td>42.97</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

***significant at 1%; **significant at 5%; *significant at 10% significance levels.

Source: Household survey (2012)
because it pays less compared to other sectors. Although there was no significant difference between the educational level of irrigators and non-irrigators, discussions with the farmers indicated that more household members of irrigators were educated compared to non-irrigators. This result, consistent with Tesfaye et al. (2008), suggests that money from irrigation is also being invested in services such as education for children.

The results show small land holdings for irrigators compared to non-irrigators implying that the intensive nature of irrigation farming is such that irrigators have to operate less land, while the non-irrigators need to put more land under cultivation to cater for the extensive and risky nature of dry-land farming. Compared to non-irrigators, the irrigators can achieve a given welfare level from a smaller land area as irrigation access enhances land productivity, ceteris paribus. Irrigators had far higher farm income compared to non-irrigators. This significant difference indicates how the climatic conditions hinder rain-fed crop production in the study area, and the farm income of non-irrigators is mainly from livestock production. Farming contributes modestly to the total household incomes in the area, contributing only about 15%. This percentage is less but close to the 18% reported by Dearlove (2007) as contribution of farming to income in the Msinga area.

Table 1 also indicates that irrigators’ welfare status is generally above that of non-irrigators, as shown by their significantly higher (by 45%) total expenditures per adult equivalent. There were significant differences between irrigators and non-irrigators in terms of both food and non-food consumption expenditure. Non-irrigators spend much of their income on food, with food expenses contributing an average of 87% of total consumption expenditure. This again shows that irrigators are better-off in terms of affluence. Women play a dominant role in both irrigation and non-irrigation farming, as shown by percentages presented in Table 2. The table below indicates that the majority (65.15%) of the households were female-headed, which supports other reports of this widely encountered phenomenon in South Africa (Denison and Manona, 2007; Mnkeni et al., 2010).

Table 2 also indicates that there was a significant difference between the support services that irrigators get compared to non-irrigators in the study area. The majority of the irrigators reported that they had access to extension, credit and good roads while the majority of the non-irrigators had none. The pressure on the government to ensure that projects do not fail results in that skewed distribution of support to irrigators. Also, irrigators perceived their soils to be of good quality, while the majority of the non-irrigators felt that their soils were infertile. Descriptive analysis presented above highlights the features of irrigators and non-irrigators. The FGT poverty analysis is presented in the next section to further characterise the irrigators and non-irrigators in terms of poverty.

**Foster Greer Thorbecke (FGT) poverty indices**

The Foster Greer Thorbecke (FGT) poverty indices indicated that the irrigators are better-off than non-irrigators. Using the poverty line of ZAR 276.64, Table 3 indicates that poverty incidence was higher among non-irrigators compared to irrigators, with 75% of non-irrigators in the study classified as poor compared to 55% of irrigators.

Poverty incidence was generally high across the whole sample, i.e., even among irrigators. The study area experiences high levels of poverty as shown by high (61%) poverty incidence. The depth and severity of poverty was higher among the non-irrigators than among irrigators. The poverty gap index is 31% for non-irrigators and 16% for irrigators. This implies that the current consumption level of poor non-irrigators and irrigators would have to increase by an average of 31% and 16%, respectively, to lift them out of poverty. The squared poverty gap index indicates that inequality among the poor is higher for non-irrigators than it is for irrigators. The FGT poverty indices indicate that although poverty is prevalent for both groups it is more pronounced among non-irrigators. The next section provides econometric models that were used to examine the extent to which irrigation increases household consumption, holding other factors constant.

**Econometric results**

The treatment effect model was used in this study to evaluate the irrigation impact on household consumption expenditure. The first stage of the treatment effect model involved estimating the participation model and the results are presented in the next section. The results of the outcome model, the second stage of the treatment effect model, are presented in the subsequent section.

**Factors predicting household’s participation in irrigation farming: Probit model results**

The binary probit model was estimated to determine the household characteristics and resource endowment that predict the household’s participation in irrigation farming. Table 4 reports the results. The results in Table 4 indicate that, collectively, all estimated coefficients are statistically significant since the LR statistic has a p-value less than 1%. The pseudo $R^2$ value is 60%, which is high for cross-sectional data. The model also correctly predicted about 91% of the cases, confirming that the model fits the data well.

The results show that factors such as land size, perceived soil fertility, household size, and access to support services (such as market and extension services) are significant predictors of irrigation participation. Accordingly, as the size of land operated by a household increases, the chances of being an irrigator decline by about 31%, other things held constant. This implies that irrigators tend to intensify their farming, while rain-fed farmers try to put more land under cultivation. This is in line with previous studies such as Tesfaye et al. (2008) and Dillon (2011).

The results also indicate those farmers who perceived their soil fertility to be good have a 21% greater chance of being irrigators than those who felt that their soils are infertile. This result is expected, as irrigation comes at a cost such that only those farmers with good land quality expecting better yield would engage in irrigation farming. This is consistent with the results from Bacha et al. (2011) and Tesfaye et al. (2008).

Increasing household size was found to increase the probability of practising irrigation farming. A unit increase in

| TABLE 3 |
|---|---|---|---|
| FGT poverty index | Non-irrigators | Irrigators | Total sample |
| Poverty headcount index (a=0) | 0.75 | 0.55 | 0.61 |
| Poverty gap index (a=1) | 0.31 | 0.16 | 0.20 |
| Squared poverty gap index (a=2) | 0.17 | 0.06 | 0.09 |

Source: Household survey (2012)
household size in adult equivalents results in an increase in the probability of irrigation participation by 7%. Increasing household size implies availability of cheap labour, hence the higher probability that the household will participate in irrigation farming. However, the relationship between household size and irrigation participation is not linear but inverted U-shaped, as shown by the negative coefficient of household size square. This implies that although increasing family size increases the likelihood of irrigation participation, after a certain family size this likelihood is reduced. Irrigation farming can only absorb a certain amount of labour, and increasing family size beyond that level reduces labour returns; hence bigger families tend to look for other opportunities that have higher returns for their labour.

The results also show that perceived access to the market increases the likelihood of households participating in irrigation farming. The farmers who reported that the market is easily accessible have 24% increased chance of being irrigators than those who reported otherwise. As irrigation is meant to enhance productivity and marketable surplus, it is not surprising that those farmers with better market access are more likely to participate in irrigation schemes. Access to extension services also improves the likelihood of irrigation participation and extension officers are more visible in irrigation projects. Farmers who reported contact with extension officers have a 12% greater chance of irrigating than those who reported otherwise. This result is consistent with findings by Gebregziabher et al. (2009).

The positive relationship between support services and irrigation can also be viewed as the tendency of government support to be concentrated on households engaged in projects, a common practice by the South African Government. Such an effort is rare in rain-fed smallholder agriculture. In Tugela Ferry, the Department of Agriculture is located right next to the irrigation scheme, making it more convenient to offer support services to the irrigators compared to the distant non-irrigators. This focus on farmers in projects has resulted in government support benefiting few people. As explained by Hall and Aliber (2010), the problem in South Africa has not been lack of support by the government, but the uneven distribution of that support to the farmers. However, it should also be highlighted that access to services by irrigators is also due to their entrepreneurship or initiatives. In reality, it is more likely that the irrigators are the enterprising farmers who take initiatives. These farmers would try to find new markets and make follow-ups to government institutions such as extension offices to get service, a practice not common among non-irrigators.
As expected, distance of farmer’s homestead from the irrigation scheme had a negative influence on the farmer being an irrigator. Proximity to the scheme implies less time taken to reach the scheme, and ensures ease of management. Since the bulk of the farmers walk to the scheme, the closer the farmer is to the scheme, the higher his/her likelihood of being an irrigator.

**The impact of access to irrigation access on household welfare: Treatment model results**

Table 5 shows the results from the second stage of the treatment effect model and OLS. Although previous studies (Tesfaye et al., 2008; Bacha et al., 2011) have highlighted the likely endogeneity between irrigation participation and consumption expenditure, the Hausman test ($F = 0.74, p = 0.39$) indicated that there was no evidence of endogeneity between the two variables at the conventional 10% significance level in this study. Thus, OLS was used in the second stage of the treatment effect model. Household distance from the scheme, although included in the selection equation, was excluded in the second stage of the treatment effect model to satisfy the condition for model identification and precise estimation of the coefficients. Although it is often difficult to find a variable that affects selection and does not affect outcome (Sartori, 2003), distance of household from the scheme on its own does not influence welfare.

The insignificant $\lambda$ in the treatment effect model indicates that there is no evidence of selection bias at the conventional 10% significance level. These findings demonstrate possible insignificant effect of unobservable factors on household consumption expenditure per adult equivalent. Therefore, the OLS results in Table 5 presented for comparison purposes are consistent and unbiased, and are close to those estimated using the treatment effect model. Henceforth, explanations concentrate on the treatment effect model results, although the same explanations do apply to the OLS results.

The treatment effect model fits the data very well, as

<table>
<thead>
<tr>
<th>Variable</th>
<th>Treatment effect model</th>
<th>OLS model</th>
</tr>
</thead>
<tbody>
<tr>
<td>Coef.</td>
<td>Std. err</td>
<td>Coef.</td>
</tr>
<tr>
<td><strong>Constant</strong></td>
<td>7 552.6</td>
<td>7383</td>
</tr>
<tr>
<td>Value of assets (Rands)</td>
<td>0.00014</td>
<td>0.00011</td>
</tr>
<tr>
<td>Off-farm income (Rands)</td>
<td>0.019**</td>
<td>0.018*</td>
</tr>
<tr>
<td>Land size (Ha)</td>
<td>722.49*</td>
<td>639.45*</td>
</tr>
<tr>
<td>Soil fertility (1=Good, 0=Poor)</td>
<td>138.78*</td>
<td>246.87</td>
</tr>
<tr>
<td>Livestock size in TLU</td>
<td>26.99</td>
<td>26.65</td>
</tr>
<tr>
<td>Age</td>
<td>37.87</td>
<td>43.76</td>
</tr>
<tr>
<td>Age square</td>
<td>0.015</td>
<td>-0.037</td>
</tr>
<tr>
<td>Gender (1=Male, 0=Female)</td>
<td>-278.01</td>
<td>-289.23</td>
</tr>
<tr>
<td>Education level (Years in school)</td>
<td>102.03**</td>
<td>100.87**</td>
</tr>
<tr>
<td>Household size in adult equivalents</td>
<td>-2 529.79***</td>
<td>-2504***</td>
</tr>
<tr>
<td>Household size square</td>
<td>148.60***</td>
<td>147.13***</td>
</tr>
<tr>
<td>Access to welfare grant (1=Yes, 0=No)</td>
<td>-358.33</td>
<td>-360.63</td>
</tr>
<tr>
<td>Access to credit (1=Yes, 0=No)</td>
<td>503.15</td>
<td>485.02</td>
</tr>
<tr>
<td>Market access (1=Yes, 0=No)</td>
<td>185.74</td>
<td>306.67</td>
</tr>
<tr>
<td>Access to extension services (1=Yes, 0=No)</td>
<td>672.34**</td>
<td>715.47**</td>
</tr>
<tr>
<td>Agricultural training (1=Yes, 0=No)</td>
<td>883.48*</td>
<td>874.44*</td>
</tr>
<tr>
<td>Access to good roads (1=Yes, 0=No)</td>
<td>1 655.41***</td>
<td>1 684.6***</td>
</tr>
<tr>
<td>Marital status (1=Married, 0=Otherwise)</td>
<td>1.427</td>
<td>3.47</td>
</tr>
<tr>
<td>Religion (1=Christian, 0=Otherwise)</td>
<td>-83.21</td>
<td>-58.49</td>
</tr>
<tr>
<td>Irrigation (1=Irrigator, 0=Non-irrigator)</td>
<td>2 216.14***</td>
<td>1 888***</td>
</tr>
</tbody>
</table>

$\rho$ -0.11858 0.16153
$\sigma$ 2 151.251 96.4707
Mills ratio ($\lambda$) -255.088 348.837
LR test of independent equations ($\rho=0$): $\chi^2$(1) 0.46
Wald $\chi^2$(39)/F-test 452.69*** 21.08***
Adj $R^2$ - 0.62
N 251 251

***Significant at 1%; **significant at 5%; *significant at 10% significance levels.
Source: Household survey (2012)
indicating the high $\chi^2$ and pseudo $R^2$ values. The Stata software program automatically corrects the standard errors (for heteroskedasticity). The estimated coefficients are, therefore, unbiased and consistent, while the standard errors are efficient.

The results indicate that irrigation access significantly increases household consumption per adult equivalent as shown by a positive estimated coefficient of irrigation access with a $p$-value less than 1%. Access to irrigation enabled farmers in the study area to practice double cropping, and to grow crops such as cabbages and potatoes commercially. As a result, the irrigators were able to generate more money and achieve higher consumption expenditure compared to their non-irrigating counter parts. The consumption expenditure per adult equivalent per year of an irrigator is ZAR2 491 more than that of a non-irrigator, ceteris paribus.

This result is consistent with the findings of previous studies (Tesfaye et al., 2008; Gebreziabher et al., 2009; Bacha et al., 2011; Kuwornu and Owusu, 2012). The conclusion here is that even though smallholder irrigation has admittedly failed in South Africa as many schemes have collapsed after government pull-out, those irrigation schemes that remain operational are playing an important role in rural poverty reduction.

The other factors that influenced household consumption were off-farm income, land size, livestock size, education level, family size and access to support services (e.g., credit, extension services, agricultural training) and infrastructure (e.g., good roads). In line with expectations, access to more land increases household welfare in the rural areas. The households are dependent upon agricultural activity for their livelihood, and more land implies better opportunities to produce more. The positive sign on livestock size also implies that having more livestock gives the households an opportunity to sell during lean periods. Some farmers, particularly those operating on drylands, reported that livestock farming was a key livelihood activity in their households.

The results also support a widely held view that education is critical in the fight against poverty. Additional years of schooling of the household head were positively related with consumption expenditure. Education implies opportunities of generating income, and implies better understanding of new and improved farming technologies. This result is consistent with findings from Tekana and Oladele (2011) and Namara et al. (2008). As expected, larger households consumed less per adult equivalent than those with fewer members. The relationship between household size and household total consumption expenditure was non-linear, indicating that as the family size increases, the welfare of a family decreases, but only up to a certain point. The U-shaped relationship between welfare and family size implies that beyond a certain point, welfare increases with increasing family size. This may be due to the labour-intensive nature of farming. Beyond a certain point, increasing labour within a household enables agricultural intensification. This result is consistent with findings from Bacha et al. (2011) and Kuwornu and Owusu (2012).

Access to support services and infrastructure (such as credit, extension, agricultural training and good road network) plays an important role in improving household welfare. Those households with better access to these services and infrastructure consumed more than those without. The farmers indicated that they use credit to buy agricultural inputs and sometimes for meeting family emergencies. Access to credit support also ensures that farmers can secure inputs in time. This leads to improved agricultural output, resulting in increased farm revenues. Extension services imply access to new technologies, which help improve agricultural production, while access to agricultural training improves farmers’ skills. Most of the farmers in the scheme use only trial and error, and those who have received some form of training are better-off as they would put these skills to use.

As expected, access to a good road network has a positive impact on household food security. Those households connected by good road networks have better opportunities than those connected with poor roads. Good road network implies ease of accessing main market centres such as Tugela Ferry town. Due to the rugged topography of the Msinga municipal area, certain areas are relatively isolated from the main centres and the farmers struggle to access potential markets and other services.

For robustness checks of the estimated irrigation impact parameter, the propensity score matching method (PSM) was used. Since there is no evidence of selection bias due to unobservables as indicated by the insignificant $\lambda$ in the treatment effect model, the PSM method would result in unbiased and robust impact estimates. The balancing property was selected in estimating propensity scores. The use of the balancing property ensures that a comparison group is constructed with observable characteristics distributed equivalently across quintiles in both the treatment and comparison groups (Smith and Todd, 2005). In constructing the matching estimates, the common support was imposed. Heckman et al. (1997) encouraged dropping treatment observations with weak common support as inferences can be made about causality only in the area of common support. All standard errors were bootstrapped with 1 000 repetitions following Smith and Todd (2005) and Dillon (2011).

Two matching methods, the nearest neighbour and Kernel matching methods, were used to estimate the impact. Comparing results across different matching methods can reveal whether the estimated project effect is robust (Khandker et al., 2010). PSM results presented in Table 6 support the conclusion that irrigation access does improve household expenditure, indicating that irrigators spend between ZAR2 170 and ZAR2 301 more than the non-irrigators depending on the matching method used.

The preceding table indicates that both the nearest neighbour and Kernel matching methods point to the fact that irrigation access has a significant effect on household welfare. The nearest neighbour matching method identified 20 matching households as control, and concluded that irrigation access

**TABLE 6**

<table>
<thead>
<tr>
<th>Matching method</th>
<th>Number of households</th>
<th>ATT</th>
<th>t-test</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Treatment</td>
<td>Control</td>
<td></td>
</tr>
<tr>
<td>Nearest neighbour</td>
<td>186</td>
<td>20</td>
<td>2 301.12 (851.75)</td>
</tr>
<tr>
<td>Kernel matching method</td>
<td>186</td>
<td>40</td>
<td>2 170.31 (612.96)</td>
</tr>
</tbody>
</table>

***Significant at 1%; **significant at 5%; *significant at 10% significance levels.
Source: Household survey (2012)
results in an increase of about ZAR 201 in consumption expenditure per adult equivalent per year over that of non-irrigators. The Kernel matching method, on the other hand, identified 40 matching households as control, and was somewhat conservative compared to the nearest neighbour matching method in calculating the impact estimate. The Kernel matching method concluded that irrigation access results in a gain of ZAR2 170 in consumption expenditure of the irrigators. The PSM, although reporting slightly lower irrigation impact estimates, supports the conclusion made by the treatment effect model that irrigation access has a significantly positive influence on consumption expenditure. This implies that the statistical results are robust.

CONCLUSION AND RECOMMENDATIONS

The objective of this study was to provide empirical evidence of the impact of smallholder irrigation on household welfare. Although the irrigators and non-irrigators had the same demographic patterns, the welfare of the irrigators was found to be better than that of non-irrigators. The analysis of the FGT poverty indices indicated that poverty is prevalent for both groups, but is more pronounced among non-irrigators. Some 75% of non-irrigators in the study area were classified as poor compared to 55% of irrigators. This study has shown that smallholder irrigation plays an important role in rural poverty reduction. The treatment effect model and propensity score matching method indicated that smallholder irrigation access plays a significantly positive role in improving household welfare. The implication of this finding is that, when operational, smallholder irrigation schemes play a huge role in poverty reduction in the rural areas of South Africa. Therefore, government investments in smallholder irrigation should continue.

However, it must be highlighted that poverty incidence is also high among irrigators, as the majority were classified as poor. This is the reason why many researchers and policy makers doubt the poverty reduction role of smallholder irrigation in South Africa. What is clear from this conclusion is that even though smallholder irrigation access reduces poverty among farmers, it is not enough on its own to eradicate poverty. This should not be interpreted as failure of smallholder irrigation, but an indication of the need for a holistic package of complementary rural development strategies where smallholder irrigation plays a part. Smallholder irrigation cannot be considered as a ‘magic bullet’ that reduces poverty on its own. Other rural micro-projects and development initiatives should be considered to complement smallholder irrigation and reduce rural poverty significantly.

REFERENCES


